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# Population size and spatial distribution of the white stork *Ciconia ciconia* in Poland in 1958 with insights into long-term trends in regional and global population

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### Abstract

Although white stork Ciconia ciconia counts go back to the XIX century, making the species a model in population investigations, substantial gaps remain in the knowledge of its large-scale population dynamics. In particular, incomplete past estimates from the core breeding areas leave the long-term changes uncertain. In this paper, we provide the earliest estimation of population size and spatial distribution of the white stork in its main stronghold (Poland) to reconstruct the species long-term trends. Based on original survey data collected during the 2<sup>nd</sup> (1958, incomplete in Poland) and 3<sup>rd</sup> (1974, successful) International White Stork Census (IWSC) we compared stork numbers in a random sample of villages surveyed on both occasions. We applied linear models to estimate the population growth rate between 1958 and 1974 and to assess spatial variation in population change across the country. Finally, we collated worldwide stork numbers obtained from all IWSCs and discuss the long-term population trends in light of new data. The stork population in Poland in 1958 was estimated at 46,100 breeding pairs and the nationwide density at 14.7 pairs/100 km<sup>2</sup>. A strong decline (by 30.2%, 1.88% per year) was noted between 1958 and 1974 across Poland with prominent spatial variation reflecting differences in local densities. The strongest declines in absolute terms affected the most abundant populations in Eastern Poland. Our data show that in the mid-20th century, the stork population in Poland was close to the current level, but it experienced a massive decline during the 1960s-1980s. This decline was consistent with trends in the worldwide and regional European populations, contrary to earlier statements indicating limited, if any, changes in the core European area. Overall, our data expand knowledge on the long-term dynamics in the white stork numbers and show that even massive changes may easily go undetected if based on non-solid data.

Keywords: Historical ecology, census, population dynamics, long-term trends, white stork, Central Europe

### Introduction

The white stork *Ciconia ciconia* (hereafter stork) is one of the few widespread, non-threatened bird species for which it is possible to make a nearly absolute assessment of the population size in areas reaching the size of countries and to track its long-term trends. Among wild animals, it is also one of the earliest-censused species. The first attempt of a large-scale inventory was made in 1876 in the former Galicia (S Poland) (Janota 1876), and inventories were then repeated many times in the entire species range. As a result, there are hundreds of census data available, carried out on various spatial and time scales. They show that the stork populations are very diverse in terms of density and population trends (Kaatz et al. 2017). The excessive amount of reliable data has also made the white stork a classical model species to investigate population fluctuations and other aspects of avian biology, thus appearing in many influential publications (e.g. Lack 1966; Perrins et al. 1991; Sæther et al. 2002; Chernetsov et al. 2006; Flack et al. 2018).

The International White Stork Census (IWSC) provides an overview of the regional and worldwide population dynamics of the species. To date, the censuses have been carried out seven times: in 1934, 1958, 1974, 1984, 1994/95, 2004/05 and

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2014 (Tryjanowski et al. 2006). Unfortunately, despite adequate methodology, thousands of participants and institutions involved as well as substantial financial resources, the national counts were sometimes incomplete, and since 1974, the global population size could only be assessed four times. Interestingly, the last IWSC in 2014 also was not completed, so the current size of the global stork population remains unknown, although an increase is evident (Thomsen et al. 2017). For the European population, the first provisional population assessment from 1958 referred to 93,000 breeding pairs (Schüz & Szijj 1960). However, it was noted that this was a minimum number rather, since it was based on uncertain or incomplete country estimates. e.g. the largest (Polish) population was assumed to be only 10,000 pairs.

Feeding conditions in both the breeding and wintering areas and on the migration routes are considered responsible for the global and regional changes in the population size of the white stork (Bairlein 1991). However, its population dynamics is more complex and depends on a multitude of factors, whose individual impacts are often difficult to disentangle. Droughts and rainfalls in the African wintering quarters (Kanyamibwa et al. 1990; Kania 2006), weather conditions in the breeding areas in early spring and during the early nestling period (Kosicki 2012; Tobolka et al. 2018), mortality during migration (Schulz 1988; van den Bossche et al. 2002), habitat alterations associated with agricultural intensity (Johst et al. 2001; Janiszewski et al. 2014), industrialization and other anthropogenic influences - are all external factors known to shape stork's population dynamics. Moreover, the strong density-dependence of the white stork (Sæther et al. 2006; Gadenne et al. 2014) also links population fluctuations with intra-specific factors, for instance age structure and competitive interactions, both affecting their productivity (Bocheński & Jerzak 2006; Kosicki & Kuźniak 2006). Data from various parts of the species range indicate that slower changes in abundance concern core areas, with dense populations, while smaller and unsaturated populations, subjected to the Allee effect, are characterized by the greater amplitude of changes in a short time (Schimkat 2004; Kaatz et al. 2017). One of the intriguing issues in stork population ecology is the unequal fluctuations of the nearby breeding populations, with seemingly similar external conditions (Tryjanowski et al. 2005; Sæther et al. 2006). The underlying mechanisms of these fluctuations often remain unrecognized; nevertheless, it seems appropriate to look for the reasons in factors that interact locally, e.g. in regional transformations of agricultural land. However, all these processes operate on a long-term perspective and were difficult to document in the past, including in the most densely populated parts of stork distribution, such as Central-Eastern Europe.

With its large area (312,679 km<sup>2</sup>), diverse habitats and regions of extensively managed farmland, Poland is of global importance for the white stork. Birds breeding in Poland are long-distance migrants, reaching their wintering grounds in East and South Africa via the easterly route along the coast of Levant. Individuals from western populations migrate across Straits of Gibraltar, although both migratory pools are not genetically distinct (Shephard et al. 2013) and contact in both breeding and overwintering areas. Ringing data and satellite tracking evidence confirm that individual white storks do show significant nest site philopatry and follow remarkably similar migration paths over many years (Berthold et al. 2004; Kania 2006). Indeed, in Poland, median dispersal distances from the natal sites are only 26 km (Chernetsov et al. 2006), yet an exchange of individuals hatched in Poland with storks from neighboring countries is not uncommon (Wuczyński 2005). An extremely distant natal dispersal concerned a bird that hatched in NE Poland and then was recorded breeding in SW France, 1897 km from the natal site (Kania 2006).

Poland has participated in each of the IWSC, but only four of them - 1974, 1984, 1994 and 2004 allowed to assess the size of the Polish stork population. They showed that, compared to other countries, Poland is inhabited by the largest number of breeding pairs, accounting for almost a quarter of the global population and three-fourths of the Central European population (Profus 2006). Northeastern Poland is also the region with the highest large-scale densities, exceeding 50 pairs/100 km<sup>2</sup> (Tomiałojć & Stawarczyk 2003). The census carried out in 1934 contained data concerning only a part (ca 120,000 km<sup>2</sup>) of the current Polish territory, which did not allow a comprehensive population assessment. Although the next census in 1958 was carried out throughout the territory of Poland and massive amounts of data were obtained, they were incomplete and not processed properly (Wuczyński et al. 2019, see below). As a consequence, the results of the 1958 census have never been published, except for a one-page note in a popular school journal (Szczepski 1968, referred also by Schüz & Szijj 1975). In contrast, the 3<sup>rd</sup> IWSC in 1974 exceeded all others in its cover and mass

participation and is considered to be the most accurate to date. The stork population in Poland was then estimated at 32,200 breeding pairs, detailed regional numbers were also given, including density in each of over 300 counties (Jakubiec 1985a). Data from 1974 turned out to be helpful for retrospective estimation of the stork population in 1958.

With data reaching back to the mid-20th century, our study fits into the definition of historical ecology, a growing, but not yet fully established academic discipline (Szabó 2015). Since historical ecology is focused on the interconnectedness of nature and human culture, the white stork constitutes an excellent example to develop this new field of scientific endeavor. Covering long-term periods of time, its demography can potentially help understanding the consequences of past historical events for the current structure, function and management of ecosystems (Egan & Howell 2005).

The first goal of this work is to estimate, for the first time, the national population of the white stork in 1958 using a sample of survey data collected during the 2<sup>nd</sup> IWSC. By doing this, we intend to shift the knowledge about the white stork abundance in the past in a key part of its geographical range backwards by 16 years. We also determine spatial variation in densities of the white stork and regional population trends between 1958 and 1974. Differentiation of trends is indicated by more recent data from long-term plots (Tryjanowski et al. 2005), as well as contemporary nationwide monitoring data (Chylarecki et al. 2018). Therefore, we assume that the changes in numbers were also not uniform within the country in the past. Finally, we collate stork numbers obtained from all consecutive IWSCs conducted in Poland and elsewhere and discuss the long-term population changes in light of our new data. It is widely accepted that stork numbers declined between the 1930s and 1980s, particularly in the west (Schulz 2004). However, changes in the eastern source population were less clear due to the lack of reliable data from the main Polish stronghold and were temporarily assumed to be stable or increasing (Bairlein 1991). We, therefore, aim to give some new insight into regional long-term trends based on quantitative data.

### Material and methods

### Methods of the censuses in 1958 and 1974

To estimate the population size of the white stork in Poland in 1958 we used the results of both the  $2^{nd}$  and  $3^{rd}$  IWSC, therefore the methods used in both

censuses are presented. The 2<sup>nd</sup> IWSC in 1958 was carried out by the questionnaire method (Wuczyński et al. 2019). The challenge was to reach the largest number of villages out of about 40,000 villages existing in Poland at that time. For this purpose, the questionnaire forms were addressed to the smallest administrative units of the time, called gromada, covering several villages. The direct recipients were teachers of rural schools, asked to provide information on the occurrence of the white storks in the subordinate school area, based on their own knowledge and interviews with pupils. The survey was preceded by press and radio census notices. In July 1958 the questionnaires were sent out to each of 8,339 gromadas, i.e. 94.9% of gromadas in Poland. Most of the questionnaires, i.e. 73.6% (6,139 gromadas) were returned by the end of the year. According to the superior administrative units, i.e. counties, questionnaires were received from 250 (77.6%) out of 322 rural counties existing in the country (Szczepski 1968).

Despite the large organizational effort and a fairly high return of questionnaires, the 1958 census failed. The reasons were large gaps in data, too complicated 4-page long questionnaire forms, their delayed distribution (during the school holiday period), and the lack of reference locations to determine the coverage of the questionnaire method. However, the underestimation of the scale of the entire undertaking was probably the main reason for the failure. With a high stork density, scarce organizational capacities of the post-war period, as well as a fledgling ornithological movement in Poland, the execution of a complete census was hardly realistic. As a result, the summary of the 2<sup>nd</sup> IWSC in Poland included only the number of counted nests, breeding success in these nests and their location (Szczepski 1968) but missed an attempt to assess the national stork population.

The experience gained in 1958 was used to improve the next census in 1974. The questionnaire method was also applied, but the forms were greatly simplified and sent in due time, at the end of June. This time, the addressees were leaders of individual villages. In addition, the survey reached them through local state administration bodies, which gave the campaign an official character (Jakubiec 1985a). 39,041 questionnaires were sent to all of 2,365 municipalities in Poland, i.e. larger administrative units that replaced gromadas, covering from one to several dozen villages. A high return rate of over 85% was achieved and only two out of 317 counties did not answer at all (Jakubiec 1985b).

Along with the questionnaire campaign, in 1974 and also 1975 stork counts were carried out in around 100 counties scattered across Poland, covering about 30% of its area. Dozens of ornithologists and amateur naturalists took part in the counts. The results from the 72 most thoroughly examined counties were compared with the questionnaire indicating its high accuracy (Jakubiec 1985b). As expected, the questionnaires underestimated the number of pairs, but the difference amounted to only 4% and was unrelated to the density of the white stork. Jakubiec (1985b) also presented a detailed confrontation of both surveys and applied indicators correcting the questionnaire results, depending on both the percentage of returned questionnaires and stork density. The final effect of the 3<sup>rd</sup> IWSC in Poland was, among others, the first estimation of the total stork population in the country, as well as the number of breeding pairs and density in each county. The results were summarized in an extensive monograph (Jakubiec 1985a). We used these solid data in our retrospective analysis.

#### Other data on stork numbers

In order to present all the available data on the longterm dynamics of the stork population in Poland, we also used data collected within the national program, Monitoring of Birds of Poland (MBP). At present, this is the only source of information about the current nationwide population size, especially due to the incomplete IWSC in 2014. MBP is a program implemented annually since 2000, resulting from obligations imposed on each EU Member State, based on a consistent methodology of field research and analysis. Specifically, the white stork is monitored under the Flagship Species Monitoring (FSM) project (Chylarecki et al. 2018), covering 40-48 randomly distributed plots in Poland, each 100 km<sup>2</sup> in size. Within each plot, all the stork nests are checked twice a season, in spring to determine nest occupancy rates and - in summer - to assess breeding success. In this paper, we used stork FSM data to show the most recent population assessment in Poland and to collate it with stork numbers in other countries.

Finally, we reviewed the scattered literature data from all successive IWSCs and showed quantitatively and graphically the changes in the world population as well as national populations of the white stork. The latter was done to compile the available information on stork numbers in 1958 in individual countries within the species range. Further, we compiled national assessments from countries belonging to two European regional sub-populations divided in line with metapopulation theory, i.e. the eastern core population, that also includes Poland, and the north-western peripheral population (Schulz 1999). For the last goal, only countries with a complete series of estimates from 1958 to 2014 were included.

### Statistical approach

### The population size in 1958

The population sizes in 1958 were backestimated using data from 28 counties chosen at random (Figure 1). Within each selected county, we only included villages that were censused both in 1958 and 1974 and where the original questionnaires have been preserved (Table S1). The counties were thus treated as a random sample of locations, where the local population size was assessed on both occasions. The total area from which the data were used in the study was 27,488 km<sup>2</sup> (range of county areas 424–1933 km<sup>2</sup>), included 784 villages (county range 8–49) and 977 nests counted in 1974 (county range 3–141, ~3% of the national population estimate in 1974) and 1251 in 1958 (county range 3–160).



Figure 1. The area of Poland divided into administrative units: voivodeships (thick lines) and counties (thin lines). Locations of 28 counties in which the white stork numbers of 1958 vs 1974 were compared are marked grey. Grey markings are adjusted to county borders from 1958.

#### Estimation of population size in 1958

We applied linear models to estimate the rate of population growth (or, rate of change) between 1958 and 1974. The log of the ratio of white stork nest numbers in 1974 and 1958 represented the response. This allowed for an easy interpretation and retrospective estimation of population size in 1958 by simply dividing the 1974 census countylevel results by the estimated rate. For example, if the number of nests in a hypothetical county was 100 and 70 in 1958 and 1974, respectively, the log of the growth rate was  $\log(70/100) = \log(0.7) = -0.357$ . Since growth rates were modelled as a function of latitude, longitude and density (see below), they were estimable for any given county. Therefore, provided the above value  $(\log(0.7) = -0.357)$  and the population in a hypothetical county in 1974, e.g. 180 pairs, the expected population size in 1958 is estimated by dividing the population size in 1974 by the growth rate (i.e.  $180/\exp(-0.357) = 180/0.7 = 257$  pairs).

### Spatial variation in population growth rate

We modelled spatial variation in the rate of change across the country by allowing clinal (W-E and S-N) trends: geographical coordinates (latitude and longitude) of the county's centre were included as predictors. Local density (number of pairs per 100 km<sup>2</sup> in 1974) was included as the third predictor to reflect county-specific habitat quality and to capture density-dependent deviations from estimated spatial trends. The global model included these three predictors and the three 2-way interactions among them. Simpler (nested) models missed one or more terms and 18 models were fitted in total. All three predictors were scaled prior to analysis to help convergence. To get valid inference from our small sample (28 counties), we used the Bayesian mode of inference with Markov chain Monte Carlo (MCMC) methods to draw samples from the parameter posterior distributions. To express the absence of prior information on model parameters, we used vague, independent priors. Specifically, we chose uniform priors within the {-10, 10} range, which induce no information on parameter estimates. Three parallel chains were simulated in WinBUGS (Spiegelhalter et al. 2003) ran in the R environment (R Core Team 2018) via R2WinBUGS library (Sturtz et al. 2005). Simulations were set conservative with 5 million samples, a burn-in of 1 million and a thinning rate of 8,000 to ensure no autocorrelation in the posterior samples and reliable inference. Chain convergence was monitored visually and by Gelman–Rubin statistics  $(\hat{R})$ : no issues were detected (all  $\hat{R}$  values were <1.005), and desired posterior sample sizes (500 draws from each chain, 1,500 samples in total) were obtained. Predictions of population size in 1958 were obtained manually from parameter posterior distributions under each model: county-specific, local densities and geographical coordinates were used as new data to produce the expected rate of change for each county. Then, population sizes known from the 1974 census were divided by the estimated rates to get the county-level population sizes in 1958. We obtained model-averaged predictions by weighting predictions computed from individual models by model weights (obtained from deviance information criterion reported by WinBUGS). Summaries were presented with medians of the posterior distributions and 95% Bayesian credible intervals (95% BCI; Gelman & Hill 2007).

#### The population size in the 21st century

The annual data from the Flagship Species Monitoring (2001–2019; Chylarecki et al. 2018) were used to estimate national stork population size during the two recent decades. Since results obtained with FSM are counts (number of breeding pairs within each square), we applied a Poisson generalized linear mixed model with a log link to estimate mean stork numbers per square  $(100 \text{ km}^2)$  for each year and each of the two density zones. The zones were predefined based on 2004 IWSC. The high-density zone covered the long-recognized strongholds for the stork in Eastern Poland and contained counties with ≥30 pairs/100 km<sup>2</sup> (total area 55 603 km<sup>2</sup>). The remaining part of the country had lower densities and represented the low-density zone (total area 250 082 km<sup>2</sup>). We again opted for Bayesian inference and fitted the model with the MCMCglmm package (Hadfield 2010) in R (R Core Team 2018). The fixed part of the model included effects of zone (two levels: high-density vs low-density) and year, while by including the plot random effect we accounted for nonindependence of data points. We again used noninformative priors for all parameters, and because the model was less demanding computationally, we simulated three shorter chains (720,000 iterations each with first 120,000 discarded as a burn-in period, a thinning rate of 600 producing 1,000 samples per chain). Convergence was satisfactory as checked visually and with Gelman-Rubin statistics  $\hat{R}$  computed with the coda package (Plummer et al. 2006; all  $\hat{R}$  values <1.002). The total population size in Poland for each year was then obtained by multiplying the posterior distributions of stork densities (mean numbers per square) in each zone by the area of a zone and summing up results for both zones.

### Results

### Population growth rate and population size estimate in 1958

All models fitted to data were successful at predicting local, county-specific population sizes in 1958 (Figure 2) and agreed that the basal (i.e. average for the whole country, with deviations captured by the three predictors and their interactions) growth rates were less than one, clearly indicating a decline between 1958 and 1974. Posterior distributions of the basal growth rates varied only a little among models, peaking around 0.7, and 95% CrI maintained roughly between 0.60 and 0.86 (Table I). This translates to a population decline of 15-40% over 16 years and an average annual growth rate,  $\lambda$ , of 0.978 (CrI: 0.968 to 0.991). In terms of absolute numbers, the strongest declines affected the highly abundant populations distributed widely across the eastern part of the country, where declines by 100-200 pairs per county were common

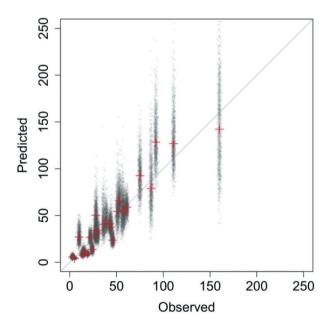


Figure 2. Predictions of the 28 county-specific population sizes in 1958. Red crosses are the medians of the model-averaged distributions of the expected population size in 1958, computed for each of the 28 counties, small translucent "x" signs are the posterior estimates (1,500 for each county). The straight line indicates a 1:1 relationship (as if model predictions matched the observed truth ideally).

(Figure 3a). The declines became smaller towards the west as most county populations in central and western Poland experienced drops by 15-70 pairs, which is not surprising, due to the smaller size of the local populations. Stable or increasing local populations, with growth rates estimated  $\geq 1$ , were distributed exclusively along the western border and rarely in the north-eastern corner (Figure 3b). It needs to be stressed that the spatial pattern of decline in absolute numbers (Figure 3a) is a product of both the spatial distribution of white stork in the country and the estimated growth rates. The strongest populations inhabit eastern Poland and become smaller towards the west, whereas the lowest estimated growth rates (meaning strongest decline relative to local population size) were noted in the area covering the southeastern corner, extending north- and westwards to reach the central part of the country (Figure 3b). So, even a low, e.g. a 0.5 rate of change (meaning population was halved, i.e. a strong decline) when applied to a small - say 20 pairs - county population gives a decline by 10 pairs in absolute numbers. A much higher rate of 0.8 applied to a more abundant population of 375 pairs would result in a decline by 75 pairs in absolute numbers. Therefore, changes in absolute numbers and growth rates do not necessarily match (Figure 3a vs b) since a very low rate may result in small declines in absolute numbers and vice versa.

The White stork population has been consistently estimated at over 40,000 pairs, depending on the model, with medians varying between 42.9 and 49.0 thousands of pairs, and average population estimates at about 46,100 pairs (Table II). It is worth noting that even the lowest limit of our model-averaged estimate (37,367 pairs) appeared higher than the estimate for 1974 (32,200 pairs) even though the confidence intervals for 1958 were not particularly narrow. Nationwide density in 1958 amounted to 14.7 pairs/100 km<sup>2</sup> and was thus higher than in 1974 (Figure 3c and d).

The Polish stork population in 2001–2019 was at a level similar to 1958. Year-to-year variations were moderate to large, with yearly means estimated between 46.2 (2007) and 61.6 (2004) thousands of pairs (Table S2). The largest year-to-year change was identified between 2004 and 2005 (Figure 4). The yearly estimates for 2001–2004 were 54.8–61.6 thousands, while in 2005–2019 they declined to 46.2–55.1 thousands of pairs. After the 2004–2005 drop, the population has not fully recovered to the levels noted prior to 2005, and a nearly linear decline was obvious over the last 6 years (Figure 4).

$\begin{array}{c c} \mbox{Model} & \mbox{Log scale} \\ \mbox{Null model} & -0.326 \\ \mbox{LON} & -0.328 \\ \mbox{LON} & -0.328 \\ \mbox{LON} & -0.325 \\ \mbox{LON+LAT} & -0.325 \\ \mbox{LON+LAT} & -0.325 \\ \mbox{LON+LAT} & -0.325 \\ \mbox{LON+LAT} & -0.327 \\ \mbox{LON+LAT} & -0.328 \\ \mbox{LON+LAT} & -0.328 \\ \mbox{LON} & -0.328 \\ \mbox{LAT} & -0.328 \\ \mbox{LAT} & -0.328 \\ \mbox{LAT} & -0.328 \\ \mbox{LON} & -0.329 \\ \mbox{LON} & -$							Interactions	
model -0.503 -0.503 -0.519 -0.519 -0.483 -0.483 -0.503 +LAT+LON×LAT -0.503 -0.503 +D -0.503 +D -0.503 -0.503 +D -0.503 -0		Normal scale	TON	LAT	D	$LON \times LAT$	TON × D	$LAT \times D$
-0.203; -0.203; -0.203; -0.483; -0.483; -0.483; -0.483; -0.503; -0.503; +D.203; -0.505; +D.203; -0.505; +D.203; -0.505; -0.505; -0.505; -0.505; -0.505; -0.505; -0.505; -0.506		0.722						
+LAT +LAT+LON×LAT +D +D +D		0.605; 0.866	-0.060					
+LAT +LAT+LON×LAT +D +D +LON×D		0.595; 0.869	-0.251; 0.135					
AT AT+LON×LAT + LON×D		0.726		0.122				
AT AT+LON×LAT + LON×D		0.614; 0.862		-0.061; 0.293				
AT+LON×LAT + LON×D		0.726	-0.061	0.123				
AT+LON×LAT + LON×D		0.602; 0.863	-0.247; 0.121	-0.072; 0.309				
Tonxd		0.721	-6.103	-1.179		6.210		
-0. -0.505; -0. -0. -0. -0. -0. -0. -0. -0.		0.604; 0.866	-9.519; -0.260	-1.957; 0.095		0.224; 9.736		
-0.505; -0. -0.508; -0. -0. -0. -0. -0. -0.		0.720			0.108			
-0. -0.508; -0. -0.504; + LON×D -0.	<b>48</b>	0.604; 0.862			-0.089; 0.297			
-0.508; -0. + LON×D -0.504;	6	0.720	-0.172		0.201			
+ LON×D	54	0.602; 0.857	-0.409; 0.055		-0.013; 0.430			
	8	0.720		0.084	0.053			
		0.604; 0.859		-0.143; 0.323	-0.174; 0.282			
	6	0.720	-0.451		-2.299		2.677	
-0.493; -0.169		0.611; 0.844	-0.790; -0.112		-4.702; 0.034		0.183; 5.246	
LAT+D+ LAT×D –0.327		0.721		0.048	-1.358			1.436
-0.508; -0.145	45	0.602; 0.865		-0.290; 0.411	-8.768; 7.129			-7.343; 8.943
LON+LAT+D –0.328	8	0.720	-0.171	0.017	0.198			
-0.504; -0.138		0.604; 0.871	-0.429; $0.068$	-0.245; 0.276	-0.093; 0.513			
LON+LAT+D+ LON×LAT	6	0.720	-5.701	-1.154	0.141	5.718		
-0.495; -0.158		0.606; 0.853	-9.635; 0.358	-2.015; 0.145	-0.139; 0.447	-0.495; 9.774		
LON+LAT+D+ LON×D -0.328	×	0.720	-0.451	0.097	-2.770		3.097	
0	l68	0.606; 0.845	-0.809; -0.106	-0.146; 0.338	-5.361; -0.096		0.281; 5.860	
LON+LAT+D+ LAT×D		0.720	-0.175	-0.025	-1.087			1.305
Ŷ		0.600; 0.858	-0.437; 0.063	-0.380; 0.345	-8.577; 7.437			-7.415; 8.954
LON+LAT+D+ LON×LAT+LON×D -0.328		0.720	-3.208	-0.520	-2.083	2.919	2.346	
-0.492; -0.162		0.611; 0.850	-9.074; 4.756	-1.864; 1.263	-5.308; 0.955	-5.475; 9.148	-0.753; 5.695	
LON+LAT+D+ LON×LAT+LAT×D -0.331		0.718	-6.198	-1.185	2.597	6.229		-2.500
Ŷ	[62	0.607; 0.850	-9.670; 0.608	-1.959; 0.184	-6.859; 9.296	-0.818; 9.785		-9.329; 7.128
LON+LAT+D+ LON×D+ LAT×D -0.331		0.718	-0.476	0.173	-0.553		3.331	-2.480
Ŷ	172	0.611; 0.842	-0.843; -0.109	-0.193; 0.513	-8.452; 5.861		0.073; 6.376	-9.342; 5.795
		0.722	-4.268	-0.617	1.405	3.992	2.577	-3.790
-0.491; -0.160		0.612; 0.852	-9.634; 3.987	-1.774; 1.171	-7.663; 7.764	-4.720; 9.617	-0.708; 5.850	-9.694; 5.264

Table I. Summary of posterior distributions of model parameters. Means and 95% CrI are given on the log scale (growth rate parameter is also shown on normal scale). The growth rate is the activated clobel rate of channel between 1058 and 1074 (see Methods). Predictor abbreviations: I.O.N. londings: I.A.T. londings: D. Lorel density in 1074. "Sources on interaction

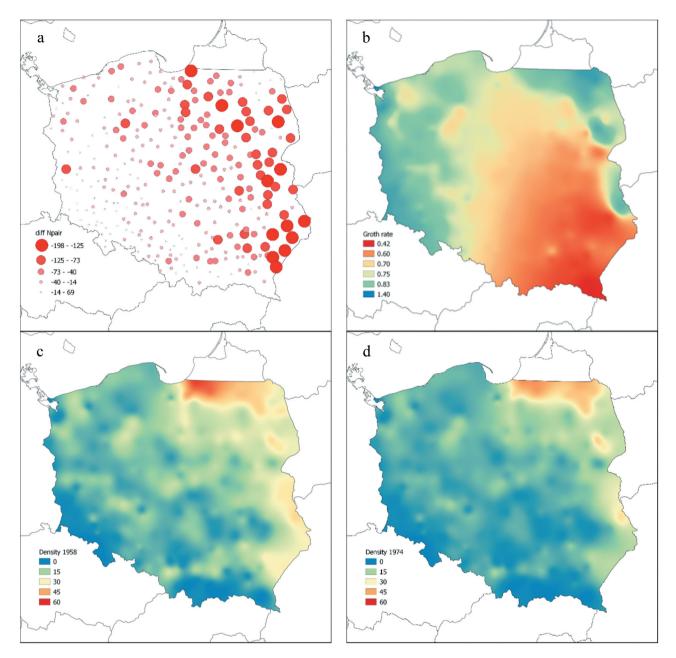


Figure 3. White stork population changes in Poland in absolute numbers (decline size in absolute numbers, pairs per county), (a), spatial distribution of growth rates (b), along with densities (pairs/100 km<sup>2</sup>) in 1958 (c) and 1974 (d).

### Discussion

### Population size in 1958 and its long-term variation in Poland

Our new estimate of the white stork population in Poland for 1958, precedes the series of later estimates and, at the same time, represents the earliest assessment for the country, allowing to follow the dynamics of the species for almost 60 years (Figure 4). In the middle of the 20th century, the population size of the white stork in Poland was high and close to the same level at the end of the century (Chodkiewicz et al. 2018; Chylarecki et al. 2018). More importantly, if this is the case, it must have undergone a massive decline since around the 1960s, up to the 1980s. This is a new finding, revealed for the first time in this study.

The size of the Polish white stork population remains unknown prior to 1958. Szczepski (1968) suggested even  $\geq 50\%$  losses between 1934 and 1958, but it was based mainly on intuition, without

		Population size estimate		
Model structure	Model weight	Median	95% BCI	
Null model	0.046	44,887	37,624–54,019	
LON	0.018	45,990	34,566-59,325	
LAT	0.033	43,161	32,977-56,010	
LON + LAT	0.018	44,598	32,247-60,998	
LON + LAT + LON×LAT	0.119	46,457	32,781-63,717	
D	0.037	42,882	33,166-55,258	
LON + D	0.048	44,032	31,829-59,554	
LAT + D	0.012	42,888	30,305-59,406	
$LON + D + LON \times D$	0.180	47,387	32,937-67,313	
$LAT + D + LAT \times D$	0.008	43,770	29,582-6,3630	
LON + LAT + D	0.016	44,130	29,888-63,532	
$LON + LAT + D + LON \times LAT$	0.081	46,207	31,219-65,806	
$LON + LAT + D + LON \times D$	0.080	47,987	31,070-72,638	
$LON + LAT + D + LAT \times D$	0.008	45,003	28,981-68,209	
$LON + LAT + D + LON \times LAT + LON \times D$	0.061	48,099	30,201-74,993	
$LON + LAT + D + LON \times LAT + LAT \times D$	0.068	45,901	29,901-6,9248	
$LON + LAT + D + LON \times D + LAT \times D$	0.063	48,979	29,748-81,407	
$LON + LAT + D + LON \times LAT + LON \times D + LAT \times D$	0.107	48,463	29,372-78,328	
Model-averaged estimate		46,129	37,367-58,280	

Table II. Models fitted to explain variation in growth rates of the White Stork population in Poland between 1958 and 1974 and the estimates of White Stork population size in 1958.

Predictor abbreviations: LON: longitude; LAT: latitude; D: density, "×" denotes an interaction term.

a sufficient basis in the data. In contrast, the only quantitative comparison based on thorough inspections made in the Milicz county in SW Poland (994  $\text{km}^2$ ) indicated stability between the 1930s

and 1950s. In 1933–1934, 119 and 152 pairs bred there, while 144 pairs were found in 1959 (Mrugasiewicz 1972). Also, Tomiałojć (1972) considered the stork population stable after evaluation

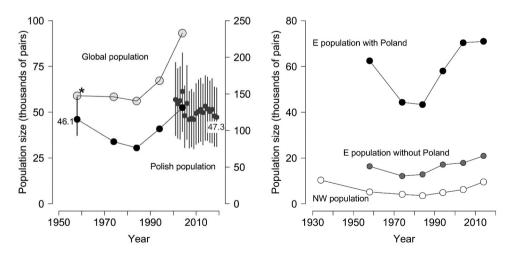


Figure 4. Changes in the number of breeding pairs in the white stork populations (see Tables III, S3 and S4 for source data). Left panel – development of the white stork populations in Poland and worldwide as obtained from consecutive international censuses 1958–2004 (black and pale grey dots) and changes of the Polish population revealed by the Monitoring of Birds of Poland (Chodkiewicz et al. 2018; Chylarecki et al. 2018) in 2001–2019 (dark grey dots). Dots show means and whiskers 95% CrI. Figures present the earliest (1958, this study) and the most recent (2019) assessments of the Polish population size in thousands of pairs. A star indicates an incomplete census. Right panel – development of the white stork (Schulz 1999) as obtained from consecutive international censuses. Six countries per each subpopulation are included, i.e. those having complete series of national assessments from the censuses in 1958–2014 and 1934 in NW population.

of incomplete field data from several areas examined both in the 1930s and 1960s; however, solid data are lacking. Nonetheless, non-linear changes in numbers cannot be ruled out during this period, i.e. significant post-war decreases, followed by rapid growth reaching the size estimated for 1958 in this work. The decrease due to World War II was likely through direct persecution of storks by soldiers (Schüz & Profus 1983), but primarily through loss of suitable agricultural habitats. Cultivation stopped on approximately 7.5 million ha of arable land in Poland and livestock was dramatically reduced: pigs and sheep by ~75%, cattle by ~70% and horses by ~50% (Bański 2010).

The strong decreasing trend between 1958 and 1974 in Poland proves rapid population changes that occurred in the most important European refuge of the white stork in the past. A 30% decline over 16 years (1.88% per year) is a fairly high figure for a long-lived, single-brooded bird species with a rather slow reproduction rate. Such fluctuations in large populations of the white stork are rare (Bairlein 1991; Sæther et al. 2006), but have already been documented. In the same period (1958 to 1974) the decline in Alsace (France) amounted to 93.1% (from 131 to 9 pairs, annual growth rate ~0.85) (Schierer 1992), in Denmark to 78.8% (from 189 to 40 pairs, annual growth rate ~0.9) (Skov 2016), whereas in Steiermark (SE Austria) the population increased by 104.1% (from 49 to 100 pairs, annual growth rate ~1.045) (Ranner & Tiefenbach 1994), and in Estonia by 121.3% (from 479 to 1060 pairs, annual growth rate ~1.051) (Ots 2009), but all these populations were much smaller than the Polish one. However, the ongoing spectacular increase of the world population of the white stork, amounting to 66% in 20 years (between 1984 and 2004, Table S3), proves the possibility of rapid population changes also at a large scale.

The causes of these various fluctuations are complex (see Introduction) and changing over time. Recently, Profus and Siekiera (2019) reported an extremely high mortality rate of first-year storks: 83% out of 110 birds equipped with satellite transmitters died within several months, mostly during their first migration. The figure is much higher than earlier estimates of the annual mortality rate (Bairlein 1991; Schaub et al. 2005). It suggests that the low survival of first-year storks may now be increasingly responsible for downward trends, as it may have been in the past. However, the twentieth-century declines were primarily linked with habitat losses and agricultural intensification. In particular, the coincidence in time between stork

declines and the mass applications of organochlorine pesticides, such as DDT (started in the 1940s, banned in the 1970s and 1980s in most developed countries, in 1976 in Poland) may not be casual. Surprisingly, this relationship was poorly emphasized, perhaps due to the lack of scientific evidence regarding the white stork itself. Further contamination studies clearly showed that various stork species (Van Den Bossche et al. 2002; Kamiński et al. 2008; Strazds et al. 2015; Orłowski et al. 2019) and other farmland birds (Pinowski et al. 1994; Orłowski et al. 2014) are susceptible to the adverse effects of agriculturally related chemicals. Therefore, it cannot be ruled out that the contamination with DDT and its breakdown products was an important cause of the past widespread population decline of white stork in Europe.

Finally, the role of intra-specific factors for the population fluctuations should not be underestimated: rather than stemming from local survival rates, productivity and immigration, declines may be due to intense dispersal to better breeding grounds, causing rapid increases there. Such examples, where a rapid growth cannot be explained without strong immigration, have already been documented for colonial waterbirds (Doxa et al. 2013; Ledwoń et al. 2014), illustrating that massive shifts of breeding range over large areas can occur. They are also exemplified by recent north- and eastward expansion of the core Central European population of the white stork, possibly facilitated by climate changes (Thomsen et al. 2017).

## Spatial variation in density and population growth rate in Poland

White stork densities in Poland in 1958 increased along an SW-NE gradient, in line with the increase of the species' main feeding grounds (pastures in particular) and areas with intensive cattle breeding. This pattern is not surprising since it closely matches spatial variation observed in later censuses (Guziak & Jakubiec 2006). In absolute numbers, the largest declines matched this distribution, and were smaller towards the west and south, along with smaller local population sizes.

However, spatial variation in the rate of change – i.e. a quantity independent of population size – between 1958 and 1974, was different. The rate of change varied primarily along an NW-SE line, perpendicular to a density gradient, which indicates that different factors than the ones affecting density variation (like habitat suitability in general) could play a role. Regional differences in the rate of change (Figure 3b) could result from processes experienced by Polish agriculture during these times. It needs to be noted that strong declines (revealed by red to yellowish areas in Figure 3b) vs stable or even increasing populations (green to blue areas on Figure 3b) match the distribution of farm ownership (private vs national) and size (small vs large) quite well. Small-sized, individual farms (which made up  $\sim 80\%$  of all in the 1950s), predominated (90–95%) in central and south-eastern Poland, i.e. in areas where the white stork population experienced strong declines. In contrast, small-sized farms of a proportion comparable with national farms were typical for areas with a more stable stork population (Kostrowicki 1978). A similar structure in farm size and ownership is still seen today (Bański 2016). This can indicate that the population of the white stork was more influenced by processes that took place in the areas of individual agricultural production with small farms (up to a few hectares), human overpopulation and an accumulation of arable lands. Even more importantly, precisely these areas saw the largest changes in land use between the 1950s and 1980s, with the increase of arable lands in particular at the cost of pastures and meadows (Kostrowicki 1978). The decline in the area of suitable feeding grounds in central and south-eastern Poland, which were already sparse and small, could affect observed regional differences in white stork population dynamics.

# Long-term trends in the Polish, global and regional stork populations

Providing the earliest estimate of the population size in Poland, we filled in the gap that made it difficult to assess the development of the white stork numbers on wider, global and regional scales. Former provisional guesses assumed that 10,000 (Schüz & Szijj 1960) or 30,000 pairs (Bairlein 1991) bred in Poland in 1958 – both are hardly defendable and turned out to be underestimated. Actually, despite incomplete national estimates, it is clear that the contribution of the Polish population to the world population was significant, constituting approx. 1/4, as in later years (Tab. S3). More importantly, our data show that trends in breeding populations in Poland did not differ from changes in the global population for over half a century. We have shown that the decline in the stork population during 1960-1980, known from other parts of Europe, and the subsequent increase, also concerned the core range of distribution covering Poland. We also

found, however, that the recent trends of the Polish and world populations diverge continuously: despite being inadequately assessed, the global trend is increasing, which is particularly prominent on the eastern and western edges of the species' range. It is however not detectable in the core range of Poland. Although the numbers are still high here (47,300 pairs in 2019 according to FSM), sharp decreases are observed in many study plots checked annually, including the key areas of Northern and Eastern Poland (Peterson & Jakubiec 2016; Sikora 2017) and are partly visible in the FSM data (Figure 4, Table S2). For example, in five provinces of S Poland (70,863 km<sup>2</sup>), a sharp decrease of 36.1% was noted between 2004 (3970 breeding pairs) and 2014 (2538 pairs) (Wuczyński et al. unpublished). During the same time, the Latvian population has increased by c 32% (Table III) and Ukrainian by c 20% (Thomsen et al. 2017) whereas the numbers in Lithuania doubled between 1994 and 2010 (Vaitkuviene & Dagys 2015). This indicates that the former center of the species range, covering Poland, is currently shifting north- and eastwards (Keller et al. 2020).

Secondly, changes in the worldwide stork population were confirmed by regional data, compiled comparatively by two European subpopulations (Figure 4). Contrary to earlier assumptions (Schulz 2004), the central-eastern population was not stable in the period of 1958–1974 but decreased at a rate of 20–30%, comparable with the peripheral north-west population. In a sample of six countries attributed to the eastern core population the rate of change during 1958-1974 was similar regardless of whether Poland was included or not (28-30%, Table II, Figure 4). Subsequent changes in both regional populations were also quite similar, up to the twenty-first century. However, the current declines in parts of Central Europe traditionally classified as the eastern core population (Poland, Czech Republic, Hungary, Austria) (Wuczyński et al. unpublished), have not kept up with the rapid growths in the north-western population (54% in 2004–2014), fed by an extremely thriving core population from the Iberia region.

### Conclusions

The white stork population dynamics in Poland, now extended by 16 years and superimposed on global and regional fluctuations, can bring a new perspective on the causes of population changes in the white stork. According to the established literature, the sharp decline of the western migratory population in

Table III. Number of breeding pairs of the White Stork obtained during consecutive IWSCs from the countries categorized into eastern core population (ECP) and north-western peripheral population (NWPP) according to Schulz (1999). Only countries with complete series of estimates from 1958 to 2014 are included, and results of the first IWSC (1934) are also presented when available. Percentage changes with respect to the previous censuses (in parentheses) along with the totals are shown at the bottom of the table. Countries are arranged according to the decreasing order of numbers in 2014. The data are also visualized in Figure 4.

	Sub-								
Country	population	1934	1958	1974	1984	1994	2004	2014	Sources
Poland	ECP		46,100	32,200	30,500	40,900	52,500	52,735	This study, Jakubiec (1985a), Jakubiec & Guziak (1998, 2006), FSM
Latvia	ECP	6750	6780	5763	6273	10,600	10,600	14,000	Janaus (2016)
Hungary	ECP		7473	4700	5100	4850	5300	4950	Profus (2006), Thomsen et al. (2017)
Slovakia	ECP	2219	1700	1124	1018	1127	1330	1351	Profus (2006), Fulin (2016)
Austria	ECP	125	276	393	319	350	395	370	Ranner & Tiefenbach (1994), Thomsen et al. (2017)
Slovenia	ECP		146	168	138	203	236	266	Profus (2006), Denac (2010), Thomsen et al. (2017)
Germany	NWPP	9035	4799	4032	3371	4135	4482	6153	Schulz (2004), Thomsen et al. (2017)
France	NWPP	155	131	18	45	315	941	2150	Schierer (1992), Profus (2006), Thomsen et al. (2017)
Netherlands	NWPP	273	56	9	8	266	563	850	Profus (2006), Thomsen et al. (2017)
Switzerland	NWPP	10	0	32	109	167	198	376	Profus (2006), Thomsen et al. (2017)
Sweden	NWPP	12	0	0	0	11	29	44	Profus (2006), Thomsen et al. (2017)
Denmark	NWPP	859	189	40	19	6	3	2	Skov (2016), Thomsen et al. (2017)
ECP withou	t Poland		16,375	12,148	12,848	17,130	17,861	20,937	
				(-25.8)	(5.8)	(33.3)	(4.3)	(17.2)	
ECP includi	ng Poland		62,475	44,348	43,348	58,030	70,361	73,672	
				(-29.0)	(-2.3)	(33.9)	(21.2)	(4.7)	
NWPP		10,344	5175	4131	3552	4900	6216	9575	
			(-50.0)	(-20.2)	(-14.0)	(38.0)	(26.9)	(54.0)	

the second half of the twentieth century was mainly associated with unfavorable changes in wintering grounds, while the presumed better state - manifested by the lower rate of decline – in the eastern population has been linked with deterioration of foraging conditions in the breeding grounds (Kanyamibwa et al. 1990). Our data suggest that changes in both - western and eastern – populations were similar to a large degree. In turn, it seems likely that the dynamics of the stork numbers were shaped by large-scale processes rather than local ones, similarly affecting birds originating from different breeding populations. Local conditions, e.g. related to agriculture, may have been responsible for the differentiation of regional trends (Senra & Alés 1992; Vaitkuviene & Dagys 2015, see also Figure 3) but unlikely for the rapid changes in the entire stork population.

Regardless of the similarities in large-scale fluctuations, we have shown variation in population changes within Poland, across the density gradient. Also in Germany, the populations in the eastern lands, adjacent to Poland, are now decreasing whereas the western and southern populations increase rapidly (Thomsen et al. 2017). This might indicate that the hitherto identification of the European subpopulations of the white stork may need to be reviewed in light of recent spatiotemporal population changes. It also means that stork censuses traditionally based on national data (such as IWSCs) may provide too coarse results, especially in large countries and regions where rapid and clinal changes occur. Although it would be rather difficult to imagine that international actions concerning numerous and widespread species could be conducted differently than at the level of individual countries, the interpretation of the results, especially regarding regional trends, should be made with caution.

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No potential conflict of interest was reported by the author(s).

### **Geolocation information**

Europe, in particular, Poland

#### Supplementary material

Supplemental data for this article can be accessed here.

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